

# Evaluation of treatment effect with paired failure times in a single arm phase 2 oncology trial

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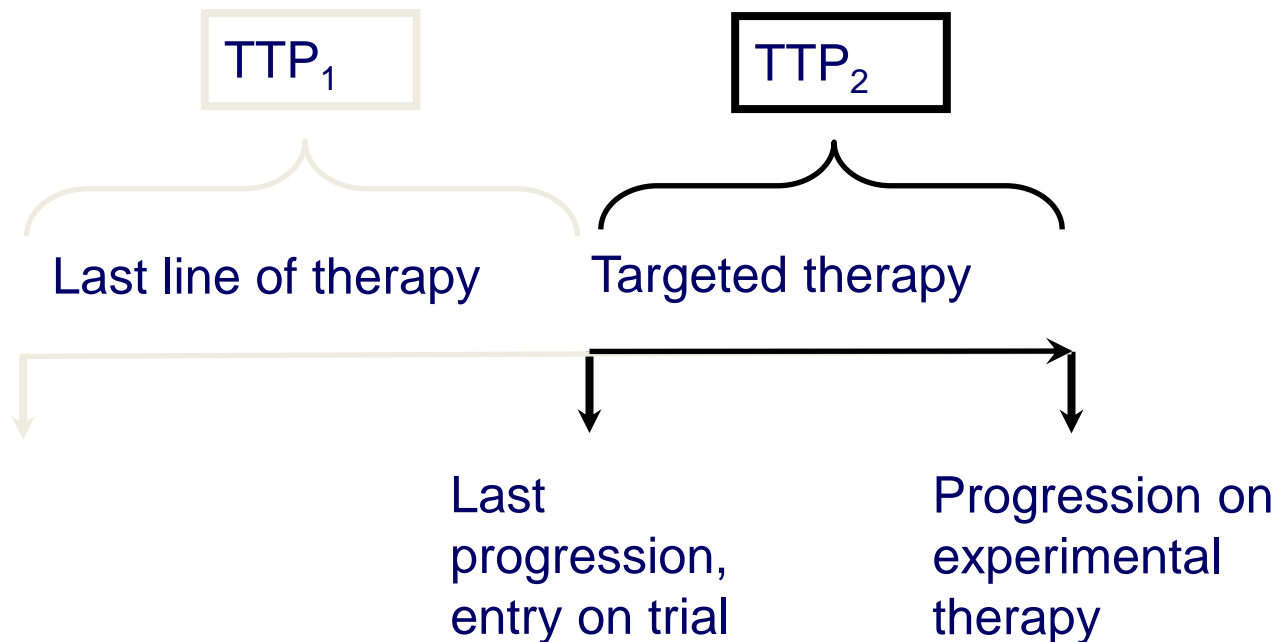


# Phase 2 Trials endpoint

- Goal of phase 2 clinical trials:  
to estimate activity and toxicity of new anticancer agents
- For cytotoxic agents,  
most used endpoints are related to the tumor shrinkage
- The problematic is different with a cytostatic agents.  
Need for an endpoint which takes account of that change

# Example of design in a multiple progression framework

- Reference *Von Hoff et al. (2010)*
- Purpose: To compare the time to progression (TTP) using the experimental treatment to the TTP with the most recent regimen on which the patient had experimented progression



# Growth Modulation index

- The growth modulation index is defined as the ratio between the time to progression of 2 successive lines of treatment:  
$$\text{GMI} = \text{TTP2} / \text{TTP1}$$
- The natural history of most advanced tumors suggests that  $\text{GMI} < 1$  (patients tend to progress increasingly faster on successive lines of treatment)
- In Von Hoff's article, the trial designed to test the hypothesis that at least 15% of the patients have  $\text{GMI} > 1.3$
- Need for correct estimate of the proportion of patients having a GMI superior to a given threshold

# Aim of our work

- To propose statistical methods to estimate the proportion of patients having their GMI superior to a given threshold by handling censored observations
- To investigate design parameters which could influence the performance of these estimators

# Methods

- We consider a study in which patients enter after having a first progression. Consequently, the time to progression at previous therapy ( $TTP_1$ ) is always observed by design

- Statistic of interest:

$$S_{GMI}(\delta) = P \left[ \frac{TTP_2}{TTP_1} > \delta \right], \quad \delta \geq 0,$$

- $\delta$  is an arbitrary threshold which represents the sign of activity considered clinically relevant

# Non-parametric approach

- The non-parametric approach, described in Kovalchik et al (2011), consists in using the ranks of each pair ( $TTP_1$ ,  $TTP_2$ ) to estimate  $S_{GMI}(1)$ . To handle censoring, we used midranks:
- $TTP_{1i}: [L_{1i}; R_{1i}]$  and  $TTP_{2i}: [L_{2i}; R_{2i}]$
- $min_{ji} : R_{j(1)} \leq R_{j(2)} \leq \dots \leq R_{j(min_i-1)} \leq L_{ji} \leq R_{j(min_i)} \leq \dots \leq R_{j(2n)}$
- $max_{ji} : L_{j(1)} \leq L_{j(2)} \leq \dots \leq L_{j(max_i)} \leq R_{ji} \leq L_{j(max_i+1)} \leq \dots \leq L_{j(2n)}$
- Imputation of the midrank:  $M_{ji} = \frac{min_{ji} + max_{ji}}{2}$
- $\hat{S}_{GMI}(1) = \frac{1}{n} \sum_{i=1}^n I(M_{2i} \geq M_{1i})$

# Parametric approach

- By assuming a parametric distribution for the GMI, the probability of interest can be derived as a function of the estimated distribution parameters
- E.g., with Weibull distributed TTPs

$$f_j(x; a, b_j | u_i) = a(u_i b_j)^{-a} x^{a-1} \exp\{-[x/(u_i b_j)]^a\}$$

the GMI has a log-logistic distribution:

$$f(\delta; a, \kappa) = a\kappa^a \delta^{a-1} (1 + (\delta\kappa)^a)^{-2}, \delta \geq 0$$

- Maximum likelihood estimates of the parameters can be obtained and used to derive the estimated probability of interest

$$S_{GMI}(\delta; \hat{\alpha}, \hat{\kappa}) = (1 + (\delta\hat{\kappa})^{\hat{\alpha}})^{-1}$$



# Simulation study

- Objectives: to evaluate the influence of the design parameters on the two estimators
- We varied:
  - The dependence between  $TTP_1$  and  $TTP_2$  (Kendall's  $\tau$ )
  - The shape of the distribution of TTP
  - The relative effect of second line treatment compared to the first
  - The censoring rate
- The statistical properties were evaluated in terms of mean bias, average standard error and empirical standard error

# Data generation

- Generation of a frailty term:  $u_i \sim \text{Gamma}$
- Generation of TTP from Weibull distribution

$$f_j(x; a, b_j | u_i) = a(u_i b_j)^{-a} x^{a-1} \exp\{-[x/(u_i b_j)]^a\}, j = 1, 2$$

with  $b_1 = e * b_2$

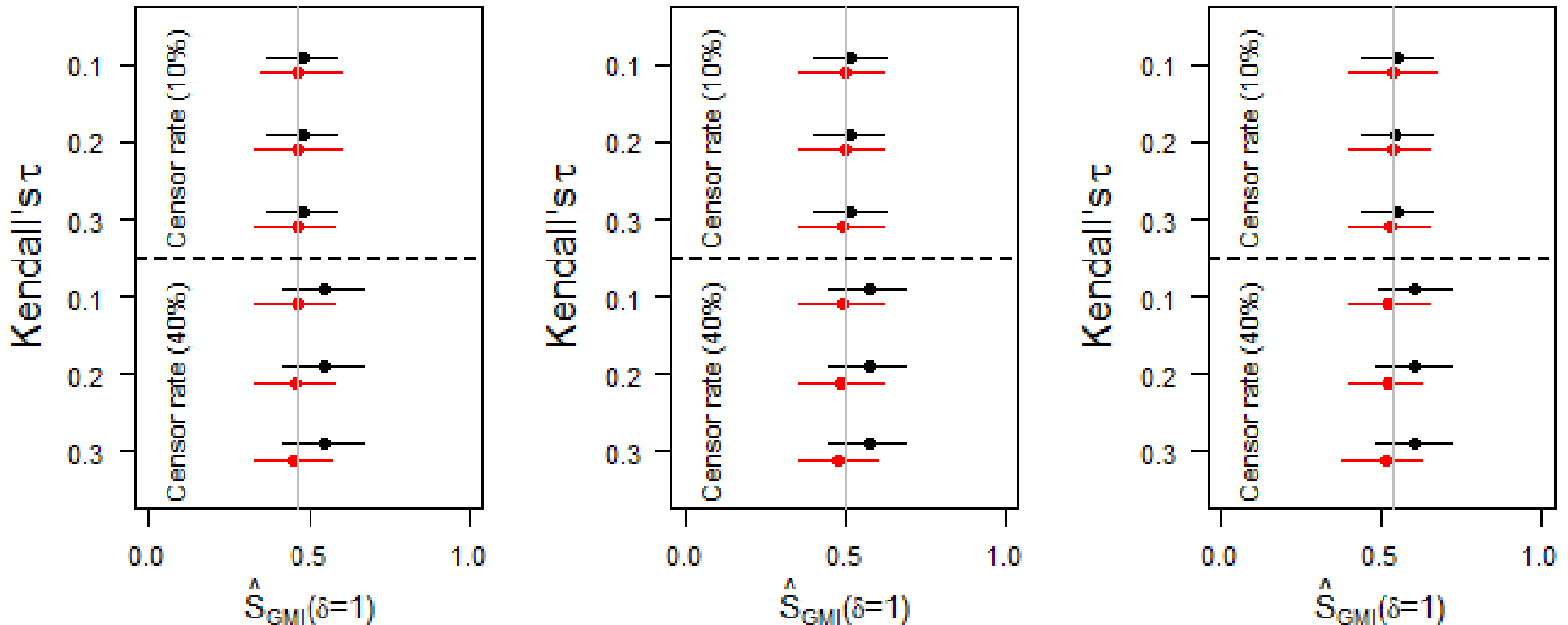
- Generation of censoring  
(10% and 40% of censored observations)

# Result of simulation (1)

med(GMI) = 0.77  
shape a = 0.5

med(GMI) = 1.00  
shape a = 0.5

med(GMI) = 1.33  
shape a = 0.5



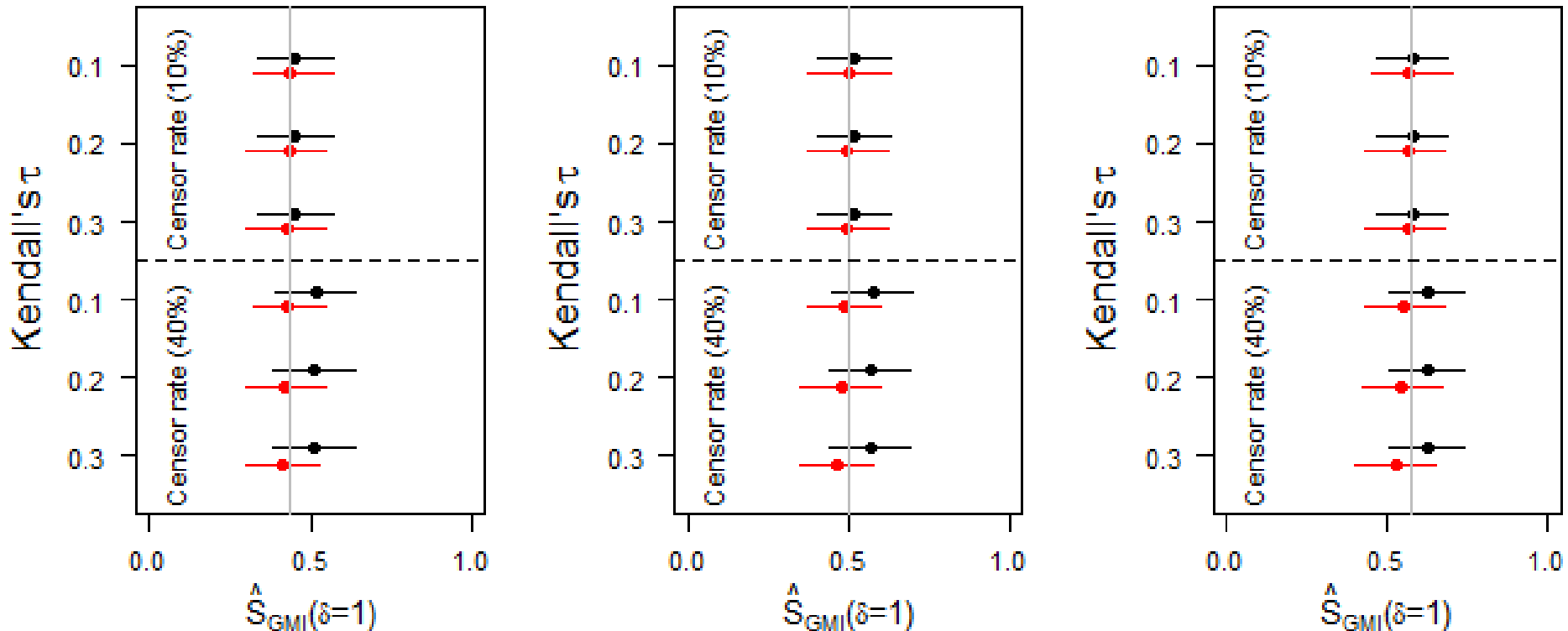
Probability  $\hat{S}_{GMI}(\delta = 1)$  of GMI being greater than 1 estimated in the simulation study via the parametric (black) and non-parametric (red) methods. Normally approximate 95% confidence intervals using the empirical standard error

# Results of simulation (2)

med(GMI) = 0.77  
shape a = 1

med(GMI) = 1.00  
shape a = 1

med(GMI) = 1.33  
shape a = 1



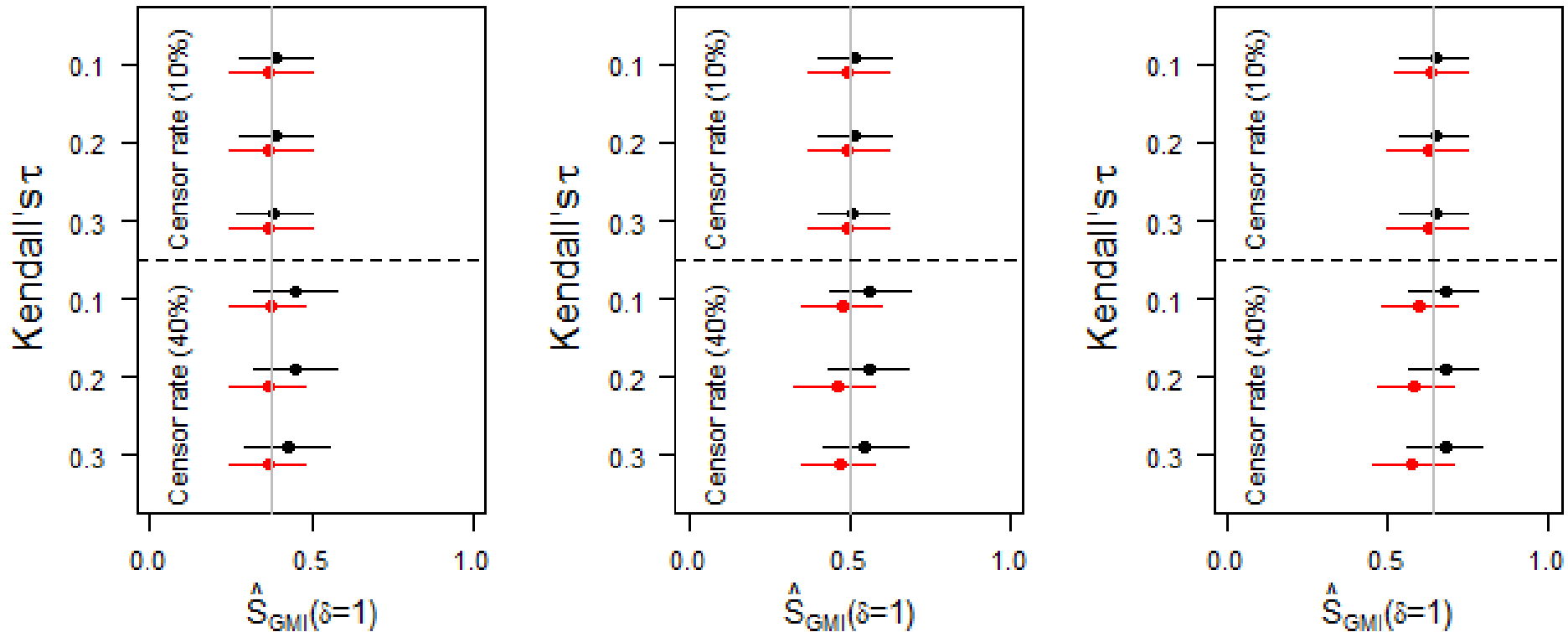
Probability  $\hat{S}_{GMI}(\delta = 1)$  of GMI being greater than 1 estimated in the simulation study via the parametric (black) and non-parametric (red) methods. Normally approximate 95% confidence intervals using the empirical standard error

# Results of simulation (3)

med(GMI) = 0.77  
shape a = 2

med(GMI) = 1.00  
shape a = 2

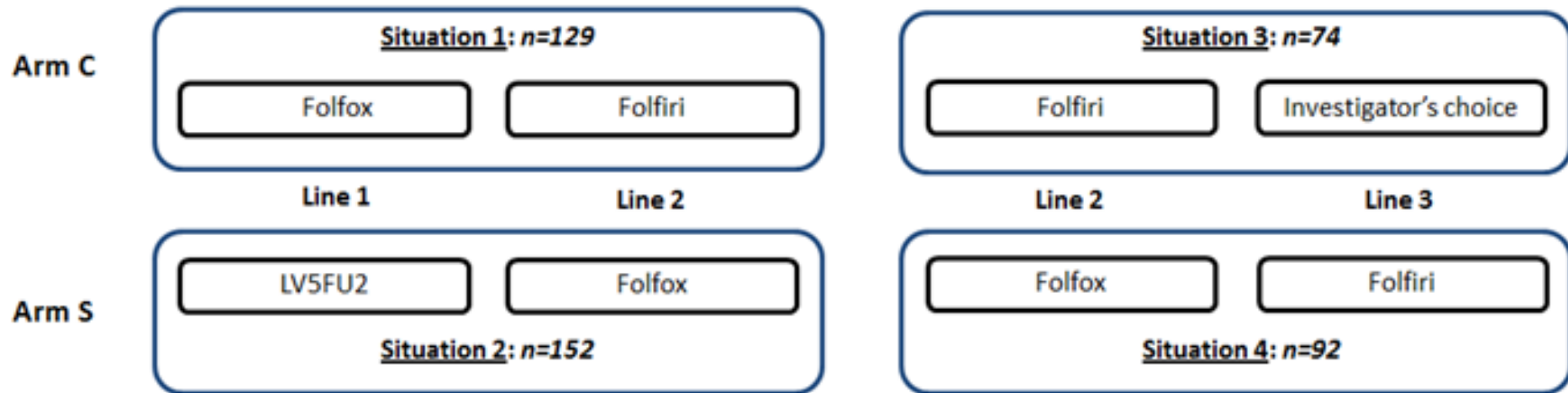
med(GMI) = 1.33  
shape a = 2



Probability  $\hat{S}_{GMI}(\delta = 1)$  of GMI being greater than 1 estimated in the simulation study via the parametric (black) and non-parametric (red) methods. Normally approximate 95% confidence intervals using the empirical standard error

# Application

- The FFCD 2000-05 trial was a randomized trial conducted by the French Federation of Digestive Oncology, which included 410 patients with advanced colorectal cancer



- Estimation of the dependence between  $TTP_1$  and  $TTP_2$  (Kendall's  $\tau$ ) by modeling the risks of progression by shared frailty models

# Dependence between $PFS_1$ and $PFS_2$

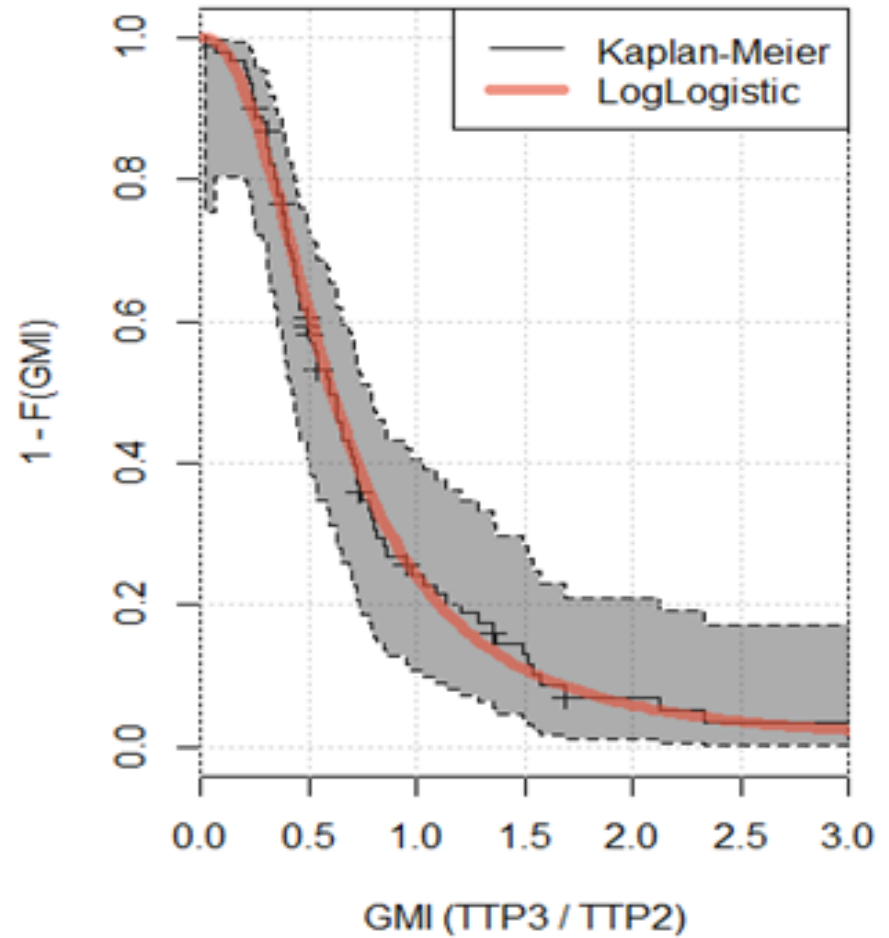
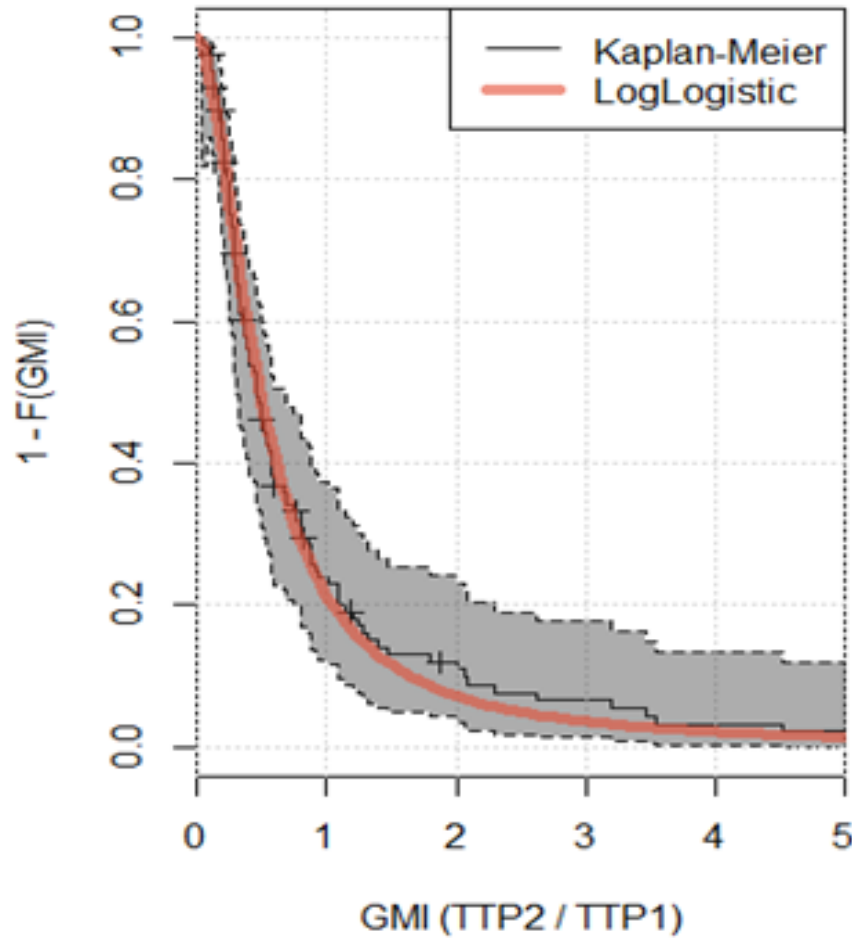
- Estimation of the Kendall's  $\tau$  for the 4 situations by modeling the risks of progression by shared frailty models

$$h_{ij}(t|u_i) = h_{0j}(t)u_i \exp(x_{ij}^T \beta)$$

- Gamma distribution assumed for the frailty term
- Weibull distribution assumed for the baseline hazard function

|              | Kendall's $\tau$ |
|--------------|------------------|
| <b>Arm C</b> |                  |
| Situation 1  | 0.195            |
| Situation 3  | 0.152            |
| <b>Arm S</b> |                  |
| Situation 2  | 0.142            |
| Situation 4  | 0.225            |

# Estimation of $\hat{S}_{GMI}(1)$



Survival function estimate of the growth modulation index (situation 1 on the left, situation 4 on the right) via the Kaplan-Meier method and via a log-logistic distribution. The gray area is the 95% confidence band for the Kaplan-Meier estimate.



# Estimation of $\hat{S}_{GMI}(1)$

|                    | Treatment |              | N   | Events | Estimator         |                   |
|--------------------|-----------|--------------|-----|--------|-------------------|-------------------|
|                    | Line 1    | Line 2       |     |        | Parametric        | Non parametric    |
| <b>Arm C</b>       |           |              |     |        |                   |                   |
| <b>Situation 1</b> | FOLFOX    | FOLFIRI      | 129 | 114    | 0.21 [0.14; 0.29] | 0.24 [0.17; 0.31] |
| <b>Situation 3</b> | FOLFIRI   | Investigator | 74  | 59     | 0.52 [0.41;0.63]  | 0.54 [0.43; 0.65] |
| <b>Arm S</b>       |           |              |     |        |                   |                   |
| <b>Situation 2</b> | LV5FU2    | FOLFOX       | 152 | 122    | 0.54 [0.46; 0.62] | 0.48 [0.40;0.56]  |
| <b>Situation 4</b> | FOLFOX    | FOLFIRI      | 92  | 79     | 0.24 [0.15; 0.33] | 0.27 [0.18; 0.36] |

Estimation of  $S_{GMI}(\delta = 1) = P(GMI > 1)$  for the four situations in the FFCD 2000-05 trial

# Discussion

- Few published clinical trials using the GMI as a criterion of activity
  - Rather low correlation of the paired time-to-progression
  - At least in some of them, this may be due to the heterogeneity of the first-line treatment or to the localization of the tumor
- In phase II trials, progressions are generally assessed at fixed times, What about the effect of interval censoring on these methods?
- An increasing number of clinical trials and the EMA admits its use to compare two successive therapies

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